



# Source-term estimation for rapid hazard assessment

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# Rapid hazard assessment

- The problem:
  - *Successful defence against harmful atmospheric releases requires timely situational awareness*
- The aim:
  - *Optimally assimilate available data to obtain a probabilistic description of the release parameters and the resulting hazard*
- The solution:
  - *Implement a real-time Monte Carlo Bayesian Data Fusion (MCBDF) algorithm for release inference*

# Bayesian data fusion

- MCBDF algorithm uses Bayesian inference over a sample set of hypothesised source-terms and Met. variables
  - it allows the set to be updated when new information is received

$$\theta = (\underbrace{x, y, t, m, a, u, v}_{\text{Source-term}}, \underbrace{L, z_0}_{\text{Met}})$$

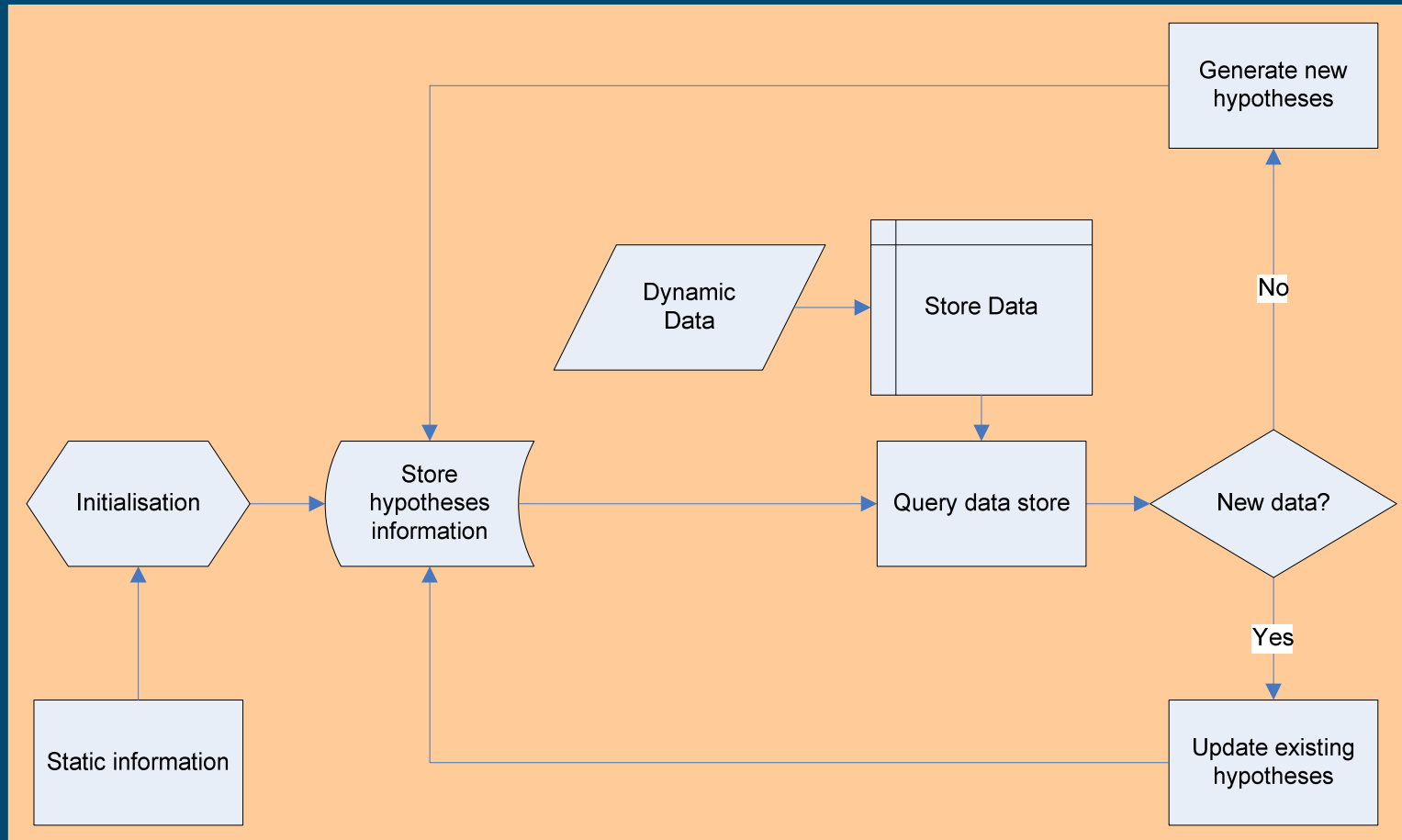
- The posterior distribution is calculated using Bayes' rule:

$$\underbrace{p(\theta|\mathbf{D})}_{\text{Posterior}} \propto \underbrace{p(\theta)}_{\text{Prior}} \underbrace{p(\mathbf{D}|\theta)}_{\text{Likelihood}}$$

# Simplistic overview of Monte-Carlo approach to source-term estimation

- Data is constantly arriving – cannot use standard Markov Chain Monte Carlo (MCMC)
- Release is a fixed point in space-time – cannot use standard Sequential Monte Carlo (SMC)
- Use a hybrid solution:
  - Fixed length, sliding data window to keep computational complexity constant
  - Hypotheses sampled and dispersion models run using MCMC in idle time
  - Hypothesis sample weights updated using stored dispersion model data when new data is received
- The combination of hypothesis weight and clustering in hypothesis space describes the *posterior* and permits probabilistic hazard assessment

# Structure of MCBDF algorithm

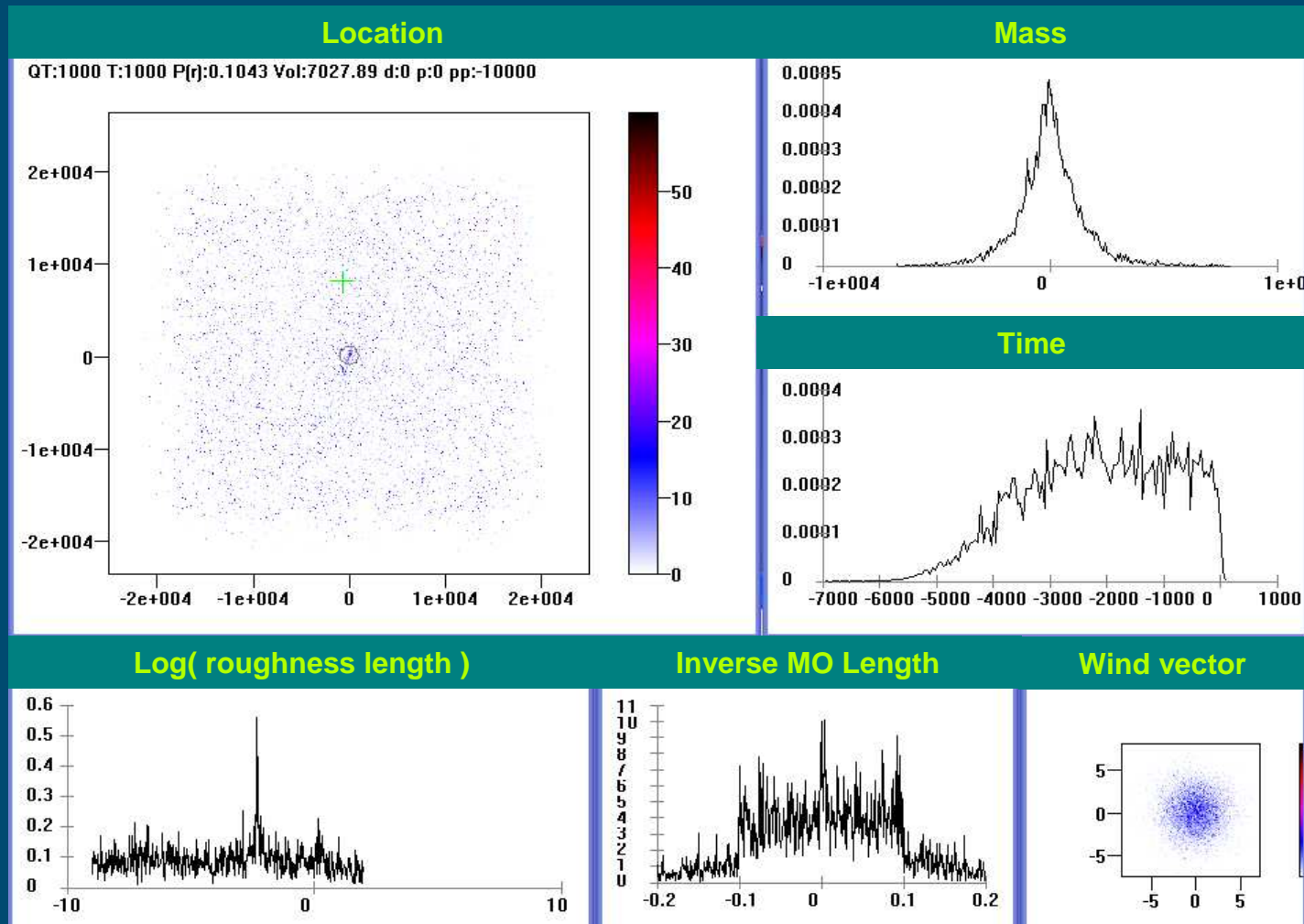


# The prior distribution

- Incorporates expert knowledge and previous experience about the shape of the hypothesis space
- Includes defining
  - Likely release locations
  - Realistic wind speeds, stabilities and surface roughness
  - Probable release times and release masses
- The mass prior is given as a double exponential

$$p(m^*) = \frac{1}{2} e^{-|m^* \mu_m|} \quad \mu_m \equiv \text{scale for release mass}$$

# MCBDF: example prior distribution



# Met Likelihood calculations

- When Met. data from a sensor is passed to MCBDF, MCBDF calculates how likely it is that the data could be generated from each hypothesis

$$\theta_k^{(i)} \equiv \text{Hypothesis } i \text{ at time slice } k$$

- The weight of each hypothesis is updated as

$$w_{k+1}^{(i)} = w_k^{(i)} p\left(d \mid \theta_k^{(i)}\right)$$



# The wind-vector likelihood model

- When wind-vector measurements are passed to MCBDF, the likelihood is calculated as

$$p\left(d \mid \theta_k^{(i)}\right) = p\left(\mathbf{u} \mid \boldsymbol{\mu}_u, \boldsymbol{\Sigma}\right) = \phi\left(\mathbf{u} \mid \boldsymbol{\mu}_u, \boldsymbol{\Sigma}\right)$$

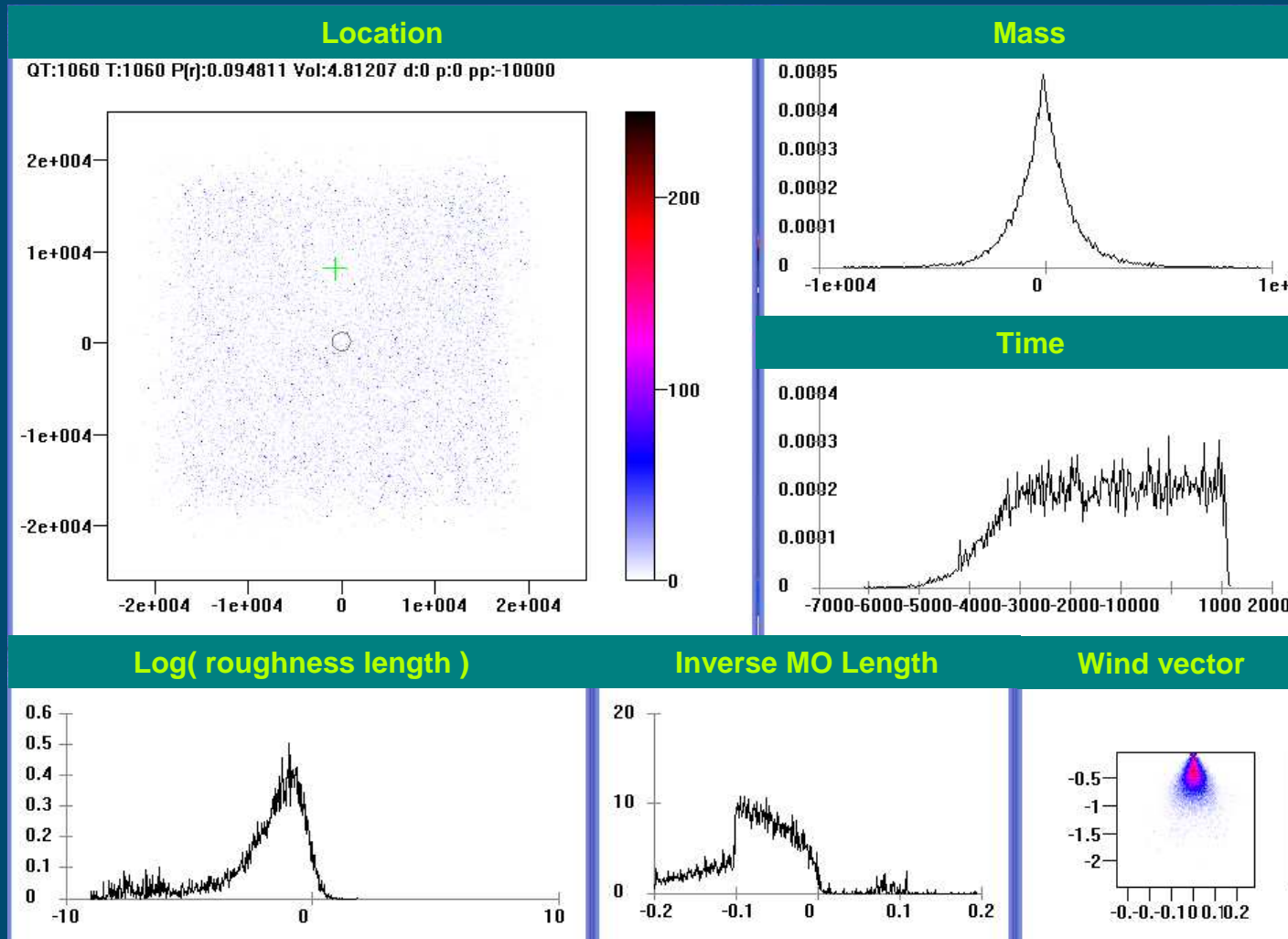
$\mathbf{u}$   $\equiv$  Measured wind-vector

 bi-variate normal distribution

$\boldsymbol{\mu}_u$   $\equiv$  Dispersion model's simulated mean wind vector at same time and location as measurement

$\boldsymbol{\Sigma}$   $\equiv$  Measurement uncertainty covariance matrix

# MCBDF: upon receipt of met data



# Detector likelihood calculations

- When CB detector measurements are passed to MCBDF the likelihood is calculated as

$$p(d | \mu, \sigma^2) = \int_0^{\infty} \underbrace{p(d | c)}_{\text{measurement density}} \underbrace{p(c | \mu, \sigma^2)}_{\text{concentration density}} dc$$

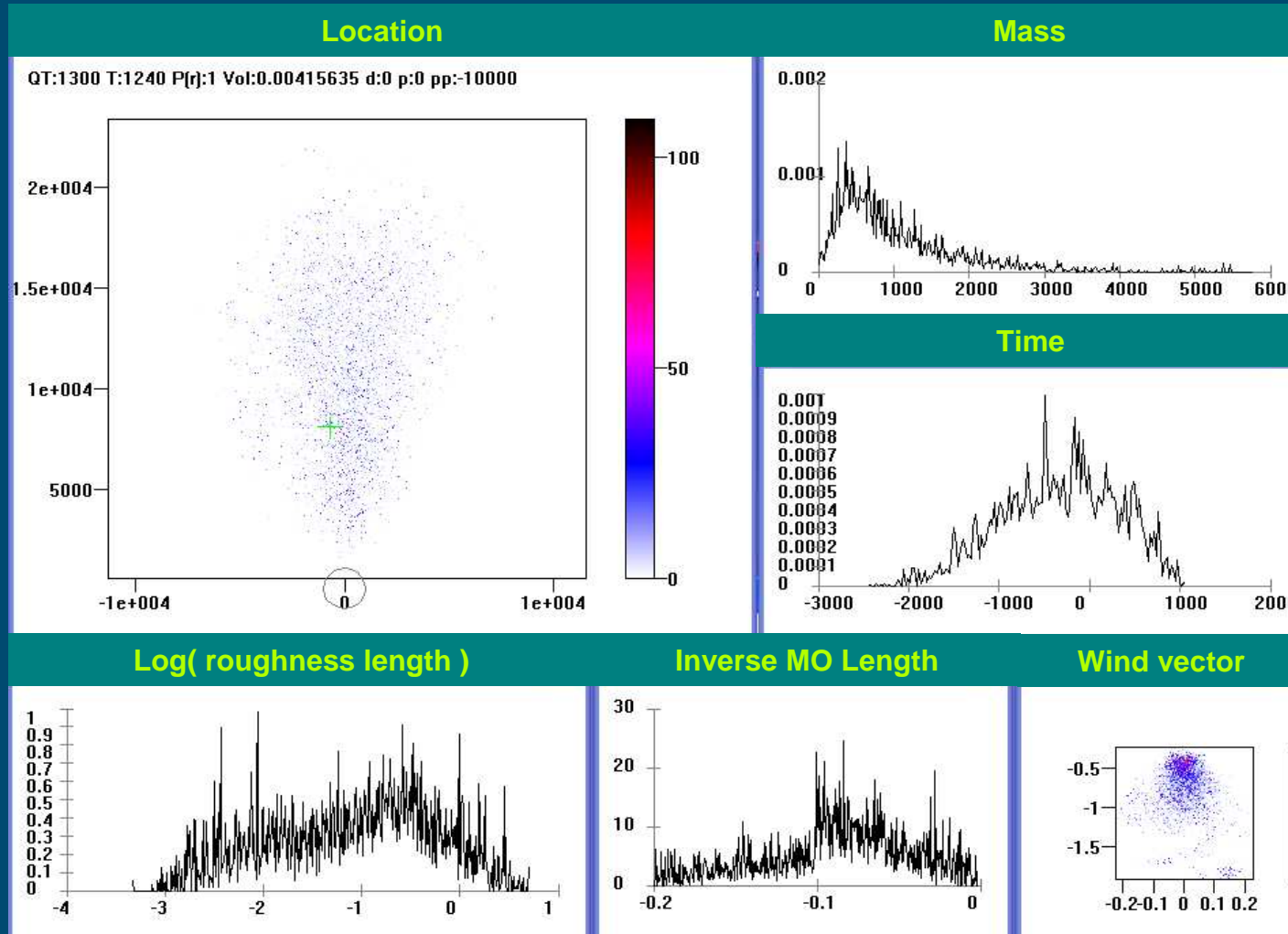
$d$   $\equiv$  Detector measurement

$\mu$   $\equiv$  Mean mass-concentration from dispersion simulation

$\sigma^2$   $\equiv$  Mass concentration variance from dispersion simulation

$c$   $\equiv$  Unobserved ground-truth concentration

# MCBDF: upon receipt of a detection



# Hypothesis generation

- MCBDF estimate of the posterior improves as it generates more and more hypotheses compatible with the data and rejects those that are incompatible
- New hypotheses are generated using Differential Evolution Markov Chain (DE-MC) Monte Carlo
- A candidate hypothesis,  $\theta_i^*$ , is generated by “mixing” up three current distinct hypotheses  $i, j$  and  $k$

$$\theta_i^* = \theta_i + \gamma(\theta_j - \theta_k) + \varepsilon$$

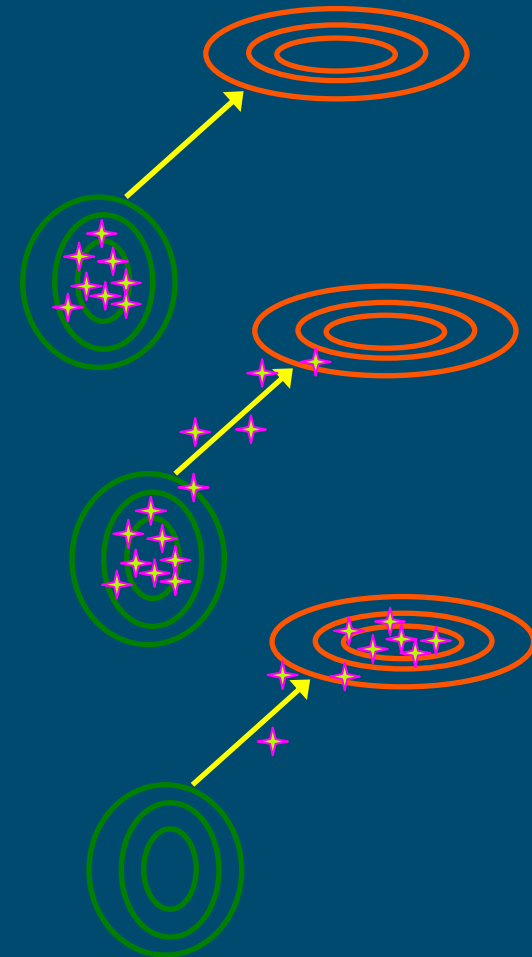
# Hypothesis generation

- The candidate hypothesis,  $\theta_i^*$ , is probabilistically added to the sample set based on the standard Metropolis accept-reject step:

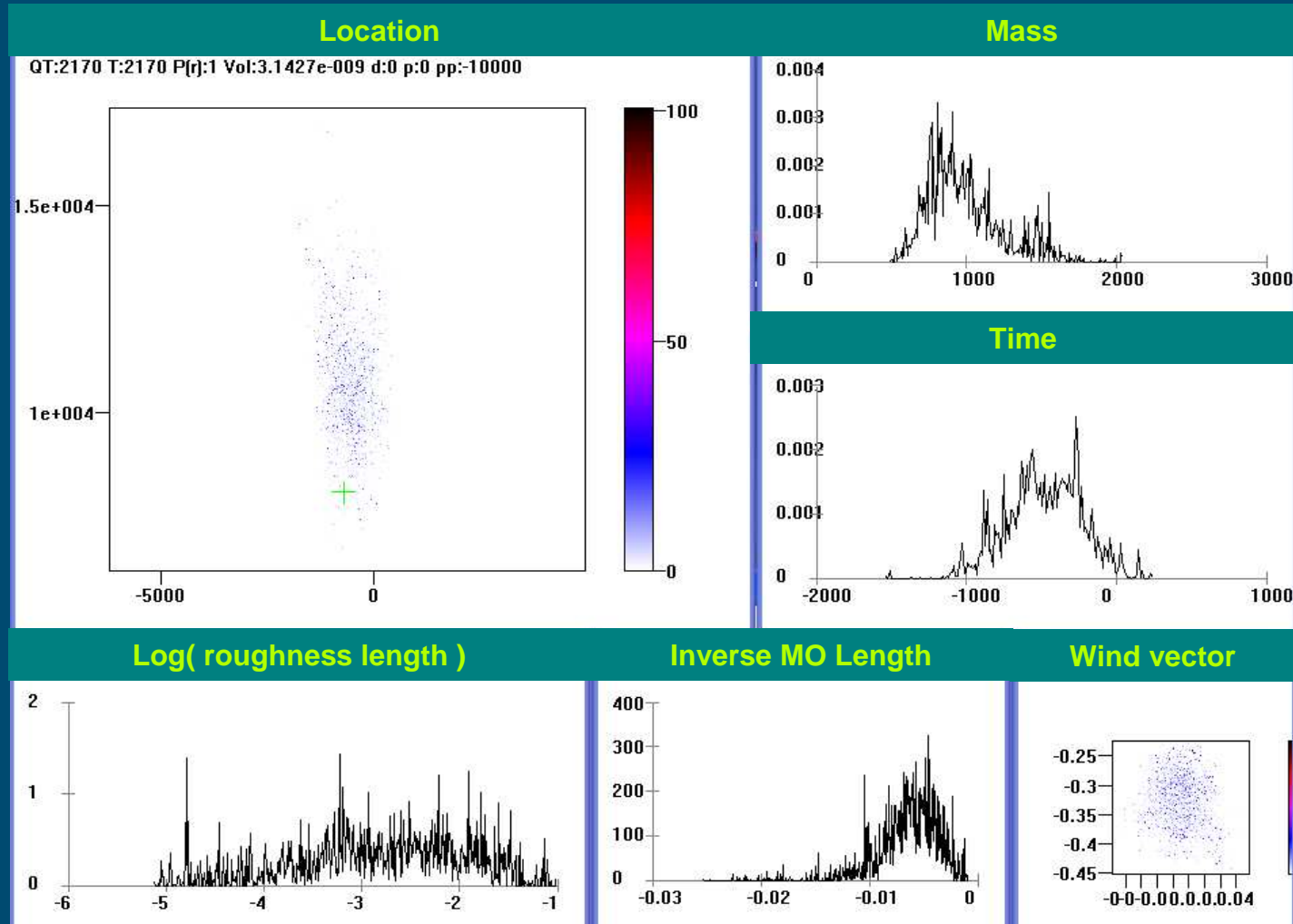
$$U(0,1) < \frac{p(\theta_i^*) \prod_{j=1}^{N_d} p(d_j | \theta_i^*)}{p(\theta_i) \prod_{j=1}^{N_d} p(d_j | \theta_i)}$$

# Burn In and Convergence

- Initial hypotheses may be far from the peak of the posterior
- But rapid answers are required
  - Incoming data continually changing posterior so population struggles to get to target
  - Once there, limited time for new sample weights to make the old ones insignificant



# MCBDF: estimate converged





# Hazard calculation

- An operational system must first determine if a hazard is present, the probability of release is calculated as

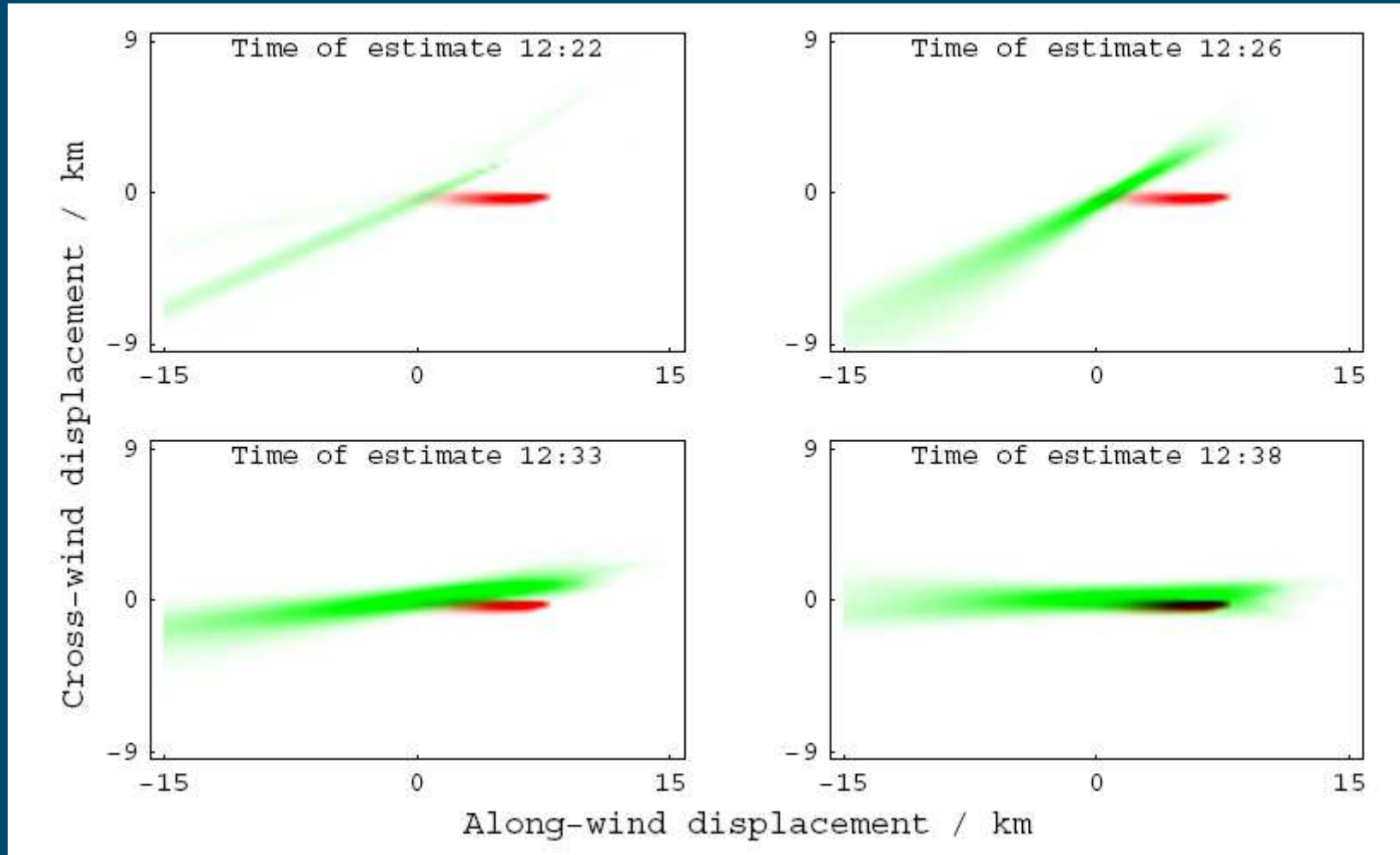
$$P(m^* > 0 | \mathbf{D}) = \frac{P(m^* > 0) \sum_i w_k^{(i)} I(m_i^* > 0)}{P(m^* > 0) \sum_i w_k^{(i)} I(m_i^* > 0) + (1 - P(m^* > 0)) \sum_i w_k^{(i)} I(m_i^* \leq 0)}$$

$w_k^{(i)}$   $\equiv$  weight of the  $i$ th hypothesis

$P(m^* > 0)$   $\equiv$  prior on a release occurring

$I(m_i^* > 0)$   $\equiv$  indicator function

# Example hazard calculation



# Summary

- Dstl have developed a real-time Bayesian inference engine capable of providing probabilistic hazard assessments
- The MCBDF algorithm can infer the uncertainty associated with the dispersion models' Met. parameters as well as release parameters
- The algorithm can therefore handle erroneous Met. input as demonstrated by its hazard calculator

# Future work

- Inference on release duration in real-time on a desktop machine
- Extension to inference on multiple releases
- Extension to inference on the dispersion model and the form of the concentration probability distribution function
- More validation with trials data
  - Demonstration with FFT07 data